



First flavor tagging calibration using 2019 Belle II data

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Abstract

We report on the first calibration of the standard Belle II B -flavor tagger using the full data set collected at the $\Upsilon(4S)$ resonance in 2019 with the Belle II detector at the SuperKEKB collider, corresponding to 8.7 fb^{-1} of integrated luminosity. The calibration is performed by reconstructing various hadronic charmed B -meson decays with flavor-specific final states. We use simulation to optimize our event selection criteria and to train the flavor tagging algorithm. We determine the tagging efficiency and the fraction of wrongly identified tag-side B candidates from a measurement of the time-integrated $B^0 - \bar{B}^0$ mixing probability. The total effective efficiency is measured to be $\varepsilon_{\text{eff}} = (33.8 \pm 3.6(\text{stat}) \pm 1.6(\text{sys}))\%$, which is in good agreement with the predictions from simulation and comparable with the best one obtained by the Belle experiment. The results show a good understanding of the detector performance and offer a basis for future calibrations.

1. INTRODUCTION AND MOTIVATION

Flavor tagging is the task of determining the heavy quark-flavor content of mesons. At Belle II, determining the flavor of neutral B mesons is needed for many measurements of $B^0 - \bar{B}^0$ mixing and CP -violation, where usually a signal B meson is fully reconstructed (signal side) and the flavor of the accompanying B -meson (tag side) has to be determined. Thus, flavor tagging plays an essential role in precise measurements of the CKM angles ϕ_1/β and ϕ_2/α and in the study of flavor anomalies that could ultimately reveal possible deviations from standard model expectations.

At Belle II, flavor tagging is accomplished using multivariate approaches. The standard algorithm is a category-based flavor tagger [1] that first identifies B^0 -decay products providing flavor information and then combines all information to determine the B^0 flavor. There is another algorithm, a deep-learning flavor tagger (DNN) [2], that determines the B^0 flavor in a single step without pre-identifying B^0 -decay products. The performance of this algorithm in Belle II data is currently being evaluated and is planned to be calibrated in the future.

In this work, we calibrate the category-based flavor tagger by measuring the time-integrated $B^0 - \bar{B}^0$ mixing probability. We reconstruct signal B decays with final states that allow us to unambiguously identify the flavor of the signal side and determine the flavor of the tag side using the flavor tagger. We reconstruct charmed signal B decays with branching fractions of 10^{-5} or larger to obtain a sufficiently large amount of signals in the current data set with a relatively straightforward reconstruction. We use the following kinematic variables to distinguish the signal from the dominant background from $e^+e^- \rightarrow q\bar{q}$ continuum events, where q indicates any quark of the first or second generation:

- the energy difference $\Delta E \equiv E_B^* - \sqrt{s}/2$ between the energy E^* of the reconstructed B candidate and half of the collision energy \sqrt{s} , both measured in the $\Upsilon(4S)$ frame;
- the beam-energy-constrained mass $M_{bc} \equiv \sqrt{s/(4c^4) - (p_B^*/c)^2}$, which is the invariant mass of the B candidate where the B energy is replaced by half the collision energy, which is more precisely known.

The signal reconstruction procedure, the event selection criteria and the training of the flavor tagger are developed and finalized using Monte Carlo (MC) simulation prior to applying it to the experimental data. Experimental and simulated data are then compared in terms of signal yields, background levels, wrong-tag fractions, tagging efficiencies and relevant distributions.

2. THE BELLE II DETECTOR

Belle II is a particle-physics detector [3, 4], designed to reconstruct the products of electron-positron collisions produced by the SuperKEKB asymmetric-energy collider [5], located at the KEK laboratory in Tsukuba, Japan. Belle II comprises several subdetectors arranged around the interaction space-point in a cylindrical geometry. The innermost subdetector is the vertex detector, which uses position-sensitive silicon layers to sample the trajectories of charged particles (tracks) in the vicinity of the interaction region to extrapolate the decay positions of their long-lived parent particles. The vertex detector includes two

inner layers of silicon pixel sensors and four outer layers of silicon microstrip sensors. The second pixel layer is currently incomplete and covers only a small portion of azimuthal angle. Charged-particle momenta and charges are measured by a large-radius, helium-ethane, small-cell central drift chamber, which also offers charged-particle-identification information through a measurement of particles' energy-loss by specific ionization. A Cherenkov-light angle and time-of-propagation detector surrounding the chamber provides charged-particle identification in the central detector volume, supplemented by proximity-focusing, aerogel, ring-imaging Cherenkov detectors in the forward regions. A CsI(Tl)-crystal electromagnetic calorimeter allows for energy measurements of electrons and photons. A solenoid surrounding the calorimeter generates a uniform axial 1.5 T magnetic field filling its inner volume. Layers of plastic scintillator and resistive-plate chambers, interspersed between the magnetic flux-return iron plates, allow for identification of K_L^0 and muons. The subdetectors most relevant for this work are the silicon vertex detector, the tracking drift chamber, the particle-identification detectors, and the electromagnetic calorimeter.

3. SELECTION AND RECONSTRUCTION OF SIGNAL B CANDIDATES

We reconstruct the following signal B decays (charge-conjugate processes are implied everywhere),

- $B^+ \rightarrow \bar{D}^0 \pi^+$,
- $B^+ \rightarrow \bar{D}^0 \rho^+$,
- $B^+ \rightarrow \bar{D}^{*0} (\rightarrow \bar{D}^0 \pi^0) \pi^+$,
- $B^+ \rightarrow \bar{D}^{*0} (\rightarrow \bar{D}^0 \pi^0) \rho^+$,
- $B^+ \rightarrow \bar{D}^{*0} (\rightarrow \bar{D}^0 \pi^0) a_1^+$,
- $B^0 \rightarrow D^- \pi^+$,
- $B^0 \rightarrow D^- \rho^+$,
- $B^0 \rightarrow D^{*-} (\rightarrow \bar{D}^0 \pi^-) \pi^+$,
- $B^0 \rightarrow D^{*-} (\rightarrow \bar{D}^0 \pi^-) \rho^+$,
- $B^0 \rightarrow D^{*-} (\rightarrow \bar{D}^0 \pi^-) a_1^+$,

for which we reconstruct the following D decays,

- $\bar{D}^0 \rightarrow K^+ \pi^-$,
- $\bar{D}^0 \rightarrow K^+ \pi^- \pi^+ \pi^-$,
- $\bar{D}^0 \rightarrow K^+ \pi^- \pi^0$,
- $\bar{D}^0 \rightarrow K_S^0 \pi^+ \pi^-$,
- $D^- \rightarrow K^+ \pi^- \pi^-$,
- $D^- \rightarrow K_S^0 \pi^-$,
- $D^- \rightarrow K_S^0 \pi^- \pi^0$,
- $D^- \rightarrow K^+ \pi^- \pi^- \pi^0$.

3.1. Data

We use generic MC simulation to optimize the event selection and compare the flavor distributions and fit results obtained from the experimental data with expectations. The generic MC simulation consists of samples that include $B^0 \bar{B}^0$, $B^+ B^-$, $u\bar{u}$, $d\bar{d}$, $c\bar{c}$, and $s\bar{s}$ processes in proportions representing their different production cross sections and correspond to an integrated luminosity of 50 fb^{-1} , about six times the $\Upsilon(4S)$ data. In addition, we generate $2 \cdot 10^7$ signal-only events [6], where the signal B meson decays to the invisible final state $B^0 \rightarrow \nu_\tau \bar{\nu}_\tau$ and the tag-side B meson decays to any possible final state according to the known branching fractions.

As for experimental data, we use all 2019 $\Upsilon(4S)$ good-quality runs, corresponding to an integrated luminosity of $8.7 \pm 0.2 \text{ fb}^{-1}$ [7]. All events are required to meet loose data-skim selection criteria, based on total energy and charged-particle multiplicity in the event, targeted at reducing sample sizes to a manageable level. All data are processed using the Belle II analysis software framework [8].

3.2. Reconstruction and baseline selection

We reconstruct charged pion and kaon candidates by starting from the most inclusive charged-particle classes and by requiring fiducial criteria that restrict them to the full acceptance in the central drift chamber and to loose ranges in impact parameter to reduce beam-background-induced tracks, which do not originate from the interaction region. Additionally, we use charged-particle identification information to identify kaon candidates. We reconstruct neutral pion candidates by requiring photons to exceed energies of about 30 MeV, restricting the diphoton mass to be in the range $120 < M(\gamma\gamma) < 145 \text{ MeV}/c^2$. The mass of the π^0 candidates is constrained to its known value in subsequent kinematic fits. For K_S^0 reconstruction, we use pairs of oppositely charged particles that originate from a common point in space or vertex position and have a dipion mass in the range $450 < M(\pi^+\pi^-) < 550 \text{ MeV}/c^2$. To reduce combinatorial background, we apply additional requirements, dependent on K_S^0 momentum, on the distance between trajectories of the two charged-pion candidates, the K_S^0 flight distance, and the angle between the pion-pair momentum and the direction of the K_S^0 flight.

The resulting K^\pm , π^\pm , π^0 , and K_S^0 candidates are combined to form $D^{(*)}$ candidates in the various final states, by requiring their invariant masses to satisfy:

- $1.84 < M(K^+\pi^-, K^+\pi^-\pi^+\pi^-, K^+\pi^-\pi^0, K_S^0\pi^+\pi^-) < 1.89 \text{ GeV}/c^2$,
- $1.844 < M(K^+\pi^-\pi^-, K_S^0\pi^-, K_S^0\pi^-\pi^0, K^+\pi^-\pi^-\pi^0) < 1.894 \text{ GeV}/c^2$,
- $0.14 < M(D^0\pi^0) - M(D^0) < 0.144 \text{ GeV}/c^2$,
- $0.143 < M(D^0\pi^+) - M(D^0) < 0.147 \text{ GeV}/c^2$.

We reconstruct ρ^\pm candidates from pairs of charged and neutral pions, and a_1^\pm candidates from three charged pions, by requiring the following conditions:

- $|M(\pi^+\pi^0) - M_\rho| < 0.1 \text{ GeV}/c^2$,
- $|M(\pi^+\pi^-\pi^+) - M_{a_1}| < 0.4 \text{ GeV}/c^2$,

where M_ρ and M_{a_1} are the known PDG masses of the ρ and a_1 mesons. To identify primary π^\pm (direct B daughters) and π^\pm candidates used to reconstruct ρ^\pm and a_1^\pm candidates, we additionally use charge-particle identification information and require the π^\pm momentum in the $\Upsilon(4S)$ frame to be larger than $0.2 \text{ GeV}/c$.

To finalize the reconstruction of signal B candidates, we associate the $D^{(*)}$ candidates with appropriate additional candidate particles π^\pm , ρ^\pm or a_1^\pm . We keep only B candidates that fulfill $M_{bc} > 5.27 \text{ GeV}/c^2$ and $|\Delta E| < 0.12 \text{ GeV}$. Additionally, for channels with ρ^\pm candidates, we remove combinatorial background from soft π^0 collinear with the ρ^\pm , by requiring

the cosine of the helicity angle θ_H between the B and the π^+ momenta in the ρ frame to be $\cos \theta_H < 0.8$.

We form the tag side of the signal B candidates, using all remaining tracks and photons that fulfill loose fiducial criteria, and KLM clusters.

3.3. Continuum suppression and final selection

To suppress continuum background from light $q\bar{q}$ pairs, we apply requirements on the two topological variables with the highest discrimination power between signal from hadronic B decays and continuum background: $\cos \theta_T^{\text{sig,tag}}$, the cosine of the angle between the thrust axis of the signal B (reconstructed) and the thrust axis of the tag-side B (remaining tracks and clusters), and R_2 , the ratio between the second and zeroth Fox-Wolfram moments using the full event information.

We vary the selection on $\cos \theta_T^{\text{sig,tag}}$ and R_2 to maximize the figure of merit $S/\sqrt{S+B}$, where S and B are the number of signal and background B candidates in the range $M_{bc} > 5.27 \text{ GeV}/c^2$ and $-0.12 < \Delta E < 0.09 \text{ GeV}$. Both $\cos \theta_T^{\text{sig,tag}}$ and R_2 requirements are optimized simultaneously using simulation. We optimize the requirements for charged and for neutral candidates independently. The optimized requirements are found to be $\cos \theta_T^{\text{sig,tag}} < 0.87$ and $R_2 < 0.43$ for charged B candidates, and $\cos \theta_T^{\text{sig,tag}} < 0.95$ and $R_2 < 0.35$ for neutral B candidates.

After applying the $\cos \theta_T^{\text{sig,tag}}$ and R_2 requirements, more than one candidate per event populates the resulting ΔE distributions, with average multiplicities for the different channels ranging from 1.00 to 7.89 (about 75% of the channels have multiplicities between 1.00 and 3.00). We select a single B candidate per event randomly to avoid possible bias using a reproducible pseudo-random ranking. The analyses of charged and neutral B channels are independent: we select one random candidate among the charged and one among the neutral channels independently.

4. THE TAGGING ALGORITHM

We determine the flavor of the tag side using the Belle II category-based flavor tagger [1]. The category-based flavor tagger is a multivariate algorithm that receives as input kinematic and PID information of the particles in the tag side, and provides as output the product $q \cdot r$, where q is the flavor of the tag-side B meson, and r the dilution factor. A dilution factor $r = 0$ corresponds to a fully diluted flavor (no possible distinction between B^0 and \bar{B}^0) and a dilution factor $r = 1$ to a perfectly tagged flavor. By convention $q = +1$ corresponds to a tag-side B^0 , and $q = -1$ to a tag-side \bar{B}^0 .

The algorithm relies on flavor-specific decay modes. Each decay mode has a particular decay topology and provides a flavor specific signature. Similar or complementary decay modes are combined to obtain additional flavor signatures. The different flavor signatures are sorted into thirteen tagging categories. Table I shows an overview of all thirteen categories together with the underlying decay modes.

The algorithm performs a two-level procedure with an event level for each category followed by a combiner level. Figure 1 illustrates the procedure. At the event level, the algorithm identifies decay products providing flavor signatures among the e^\pm , μ^\pm , K^\pm , π^\pm and Λ candidates in the tag side using Fast Boosted Decision Tree (FBDT) [9] classifiers.

At the combiner level, the algorithm combines the information provided by all categories into the final product $q \cdot r$ using a combiner-level FBDT. This classifier receives an input from each category corresponding to the product $q_{\text{cand}} \cdot y_{\text{cat}}$, where q_{cand} is the charge of the candidate identified as flavor-specific decay product, and y_{cat} is the probability provided by the event-level FBDT. Only for the Kaon and the Lambda category, the input is the effective product $(q_{\text{cand}} \cdot y_{\text{cat}})_{\text{eff}}$ of the three candidates with the highest probability.

The algorithm is trained using signal MC events where the signal B meson decays to the invisible final state $B^0 \rightarrow \nu_\tau \bar{\nu}_\tau$. Using the $B^0 \rightarrow \nu_\tau \bar{\nu}_\tau$ samples, we avoid possible bias due to CP asymmetries or reconstruction performance since these samples are generated without built-in CP violation, and all reconstructed objects (tracks, photons and KLM clusters) can be used to form the tag side without passing through reconstruction of the signal side. The flavor tagger is trained with a sample of about 10^7 MC events and tested with an independent sample of the same size to exclude overtraining.

TABLE I. Tagging categories and their targets (left) with examples of the considered decay modes (right). Here, p^* stands for momentum in the center-of-mass frame and ℓ^\pm for charged leptons (μ^- or e^-).

Categories	Targets for \bar{B}^0	Underlying decay modes
Electron	e^-	
Intermediate Electron	e^+	$\bar{B}^0 \rightarrow D^{*+} \bar{\nu}_\ell \ell^-$
Muon	μ^-	$\quad \quad \quad \downarrow D^0 \pi^+$
Intermediate Muon	μ^+	$\quad \quad \quad \quad \quad \downarrow X K^-$
Kinetic Lepton	ℓ^-	
Intermediate Kinetic Lepton	ℓ^+	$\bar{B}^0 \rightarrow D^+ \pi^- (K^-)$
Kaon	K^-	$\quad \quad \quad \downarrow K^0 \nu_\ell \ell^+$
Kaon-Pion	K^-, π^+	
Slow Pion	π^+	
Maximum p^*	ℓ^-, π^-	$\bar{B}^0 \rightarrow \Lambda_c^+ X^-$
Fast-Slow-Correlated (FSC)	ℓ^-, π^+	$\quad \quad \quad \downarrow \Lambda \pi^+$
Fast Hadron	π^-, K^-	$\quad \quad \quad \quad \quad \downarrow p \pi^-$
Lambda	Λ	

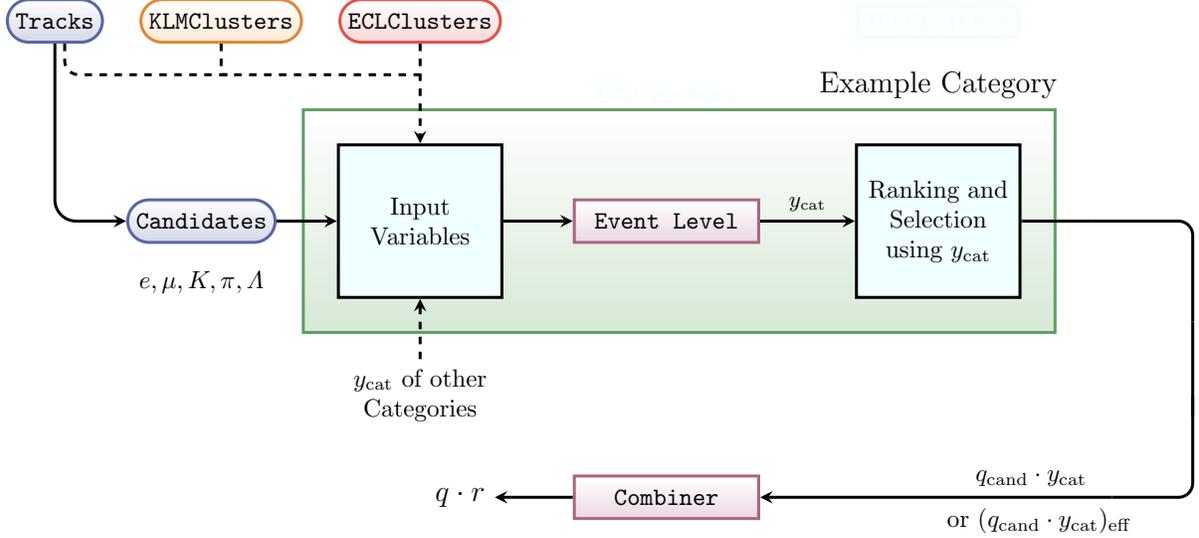


FIG. 1. Procedure for each single category (green box): the candidates correspond to the reconstructed tracks for a specific mass hypothesis. Some of the input variables consider all reconstructed tracks and all neutral ECL and KLM clusters on the tag side. The magenta boxes represent multivariate methods: y_{cat} is the output of the event level. The output of the combiner is equivalent to the product $q \cdot r$.

5. DETERMINATION OF EFFICIENCIES AND WRONG-TAG FRACTIONS

The tagging efficiency of the flavor tagger corresponds to the fraction of events to which a flavor tag can be assigned. Since the algorithm needs only one charged track on the tag side to provide a tag, the tagging efficiency is close to 100%, with good consistency between data and simulation as Table II shows.

TABLE II. Tagging efficiencies $\varepsilon \pm \delta\varepsilon$ for charged and neutral $B \rightarrow D^{(*)}h^+$ candidates in data and in simulation. All values are given in percent. The uncertainties are only statistical.

Channel	MC	Data
$B^0 \rightarrow D^{(*)-}h^+$	99.78 ± 0.02	99.78 ± 0.04
$B^+ \rightarrow \bar{D}^{(*)0}h^+$	99.81 ± 0.01	99.72 ± 0.04

To measure the fraction of wrongly tagged events w , we sort the events in bins of the dilution factor r provided by the flavor tagger. To compare with our predecessor experiment, we use the binning introduced by Belle [10].

Considering $\mathcal{T}(4S) \rightarrow B^0 \bar{B}^0$ events, the time-integrated probability to observe an event with signal B flavor $q_{\text{sig}} \in \{-1, +1\}$ and tag-side B flavor $q_{\text{tag}} \in \{-1, +1\}$ in the i -th r bin, is given by

$$\mathcal{P}^i(q_{\text{sig}}, q_{\text{tag}}) = \frac{1}{2} \varepsilon_i \left[1 - q_{\text{sig}} \cdot q_{\text{tag}} \cdot (1 - 2w_i) \cdot (1 - 2 \cdot \chi_d) \right], \quad (1)$$

where χ_d is the $B^0 - \bar{B}^0$ mixing probability, and w_i and ε_i are the wrong-tag fraction and the partial tagging efficiency in the i -th r bin (7 bins in total). The expression above is obtained

assuming that the signal B flavor is correctly identified and that there is no asymmetry in the performance between B^0 and \bar{B}^0 events. We neglect those possible small asymmetries due to the small size of the currently available data sample. The current world average for the $B^0 - \bar{B}^0$ mixing probability is $\chi_d = 0.1858 \pm 0.0011$ [11].

Since we need to consider the background to determine the signal w_i and ε_i , we developed a statistical model with a signal and a background component. We determine the signal yield N_{sig} , the background yield N_{bkg} , the partial efficiencies ε_i and the wrong-tag fractions w_i from an extended maximum likelihood fit to the unbinned distributions of ΔE , q_{sig} and q_{tag} . We checked that the ΔE distribution is statistically independent from those of q_{sig} and q_{tag} with Pearson correlation coefficients below 2%.

In the fit model, the probability density function (PDF) for each component j is given by

$$\mathcal{P}_j(\Delta E, q_{\text{sig}}, q_{\text{tag}}) \equiv \mathcal{P}_j(\Delta E) \cdot \mathcal{P}_j(q_{\text{sig}}, q_{\text{tag}}).$$

We model the signal ΔE PDF using a Gaussian plus a Crystal Ball function [12] determined empirically using correctly associated signal MC events, with the additional flexibility of a global shift of peak position and a global scaling factor for the width as suggested by a likelihood-ratio test. The background ΔE PDF is modeled using an exponential function with a free-to-float exponent.

The flavor PDF $\mathcal{P}(q_{\text{sig}}, q_{\text{tag}})$ has the same form for signal and background (Eq. 1) with independent ε_i , w_i and χ_d parameters for signal and background. We fix the background χ_d^{bkg} parameter to 0 as we obtain values compatible with 0 when we let it float.

The total extended likelihood is given by

$$\mathcal{L} \equiv \prod_i \frac{e^{-\sum_j N_j \varepsilon_i}}{N^i!} \prod_{k=1}^{N^i} \sum_j N_j \cdot \mathcal{P}_j^i(\Delta E^k, q_{\text{sig}}^k, q_{\text{tag}}^k),$$

where i extends over the r bins, k extends over the events in the bin i , and j over the two components: signal and background. The PDFs for the different components have no common parameters. N^i denotes the total number of events in the i -th r bin. The partial efficiencies ε_i are included in the flavor part of \mathcal{P}_j . Since we can fit only to events with flavor information, the sum of all ε_i must be 1. We therefore replace the epsilon for the first bin (with lowest r) with

$$\varepsilon_1 = 1 - \sum_{i=2}^7 \varepsilon_i, \tag{2}$$

and obtain its uncertainty $\delta\varepsilon_1$ from the width of the residuals of pseudo-experiments.

To validate the ΔE model, we first perform an extended maximum likelihood fit to the unbinned distribution of ΔE (without flavor part) in simulation and data. Figures 2 and 3 show the ΔE fit projections in data and simulation for charged and neutral $B \rightarrow D^{(*)}h^+$ candidates. Table III summarizes the yields obtained from the fits. We observe a good agreement between data and simulation for neutral B candidates, and lower signal yield with respect to the expectation for charged B candidates.

To determine the partial efficiencies ε_i and the wrong-tag fractions w_i , we perform a fit of the full model in a single step. For neutral candidates, we additionally leave the signal χ_d^{sig} free to float constraining it via a Gaussian constraint,

$$\mathcal{L} \Rightarrow G(\chi_d^{\text{sig}} - \chi_d, \delta\chi_d) \cdot \mathcal{L},$$

where χ_d and $\delta\chi_d$ are the central value and the uncertainty of the world average. For charged B mesons, χ_d is equal to 0 as there is no flavor mixing due to electric charge conservation.

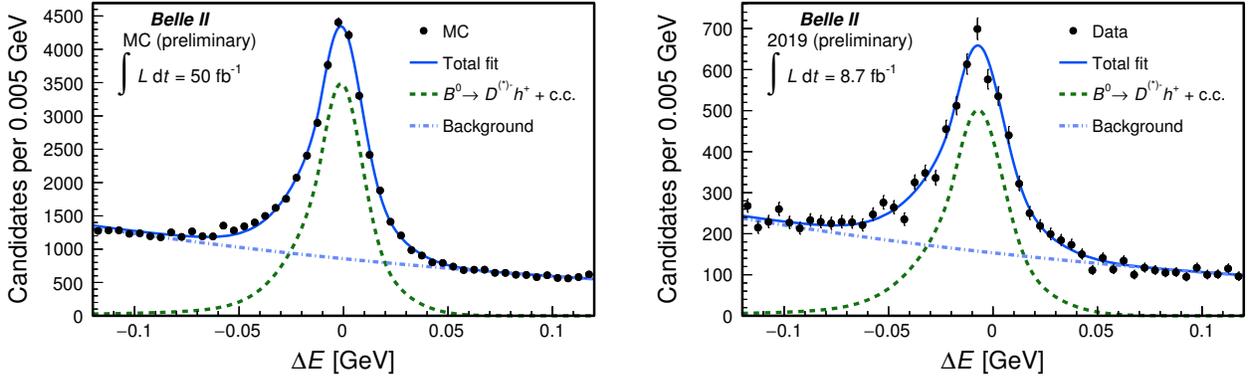


FIG. 2. Fit projection of the maximum likelihood fit to the unbinned distribution of ΔE for $B^0 \rightarrow D^{(*)-} h^+$ candidates reconstructed in (left) simulation and (right) data, restricted to $M_{bc} > 5.27 \text{ GeV}/c^2$. The global peak shift and width scaling factor are determined by the fit.

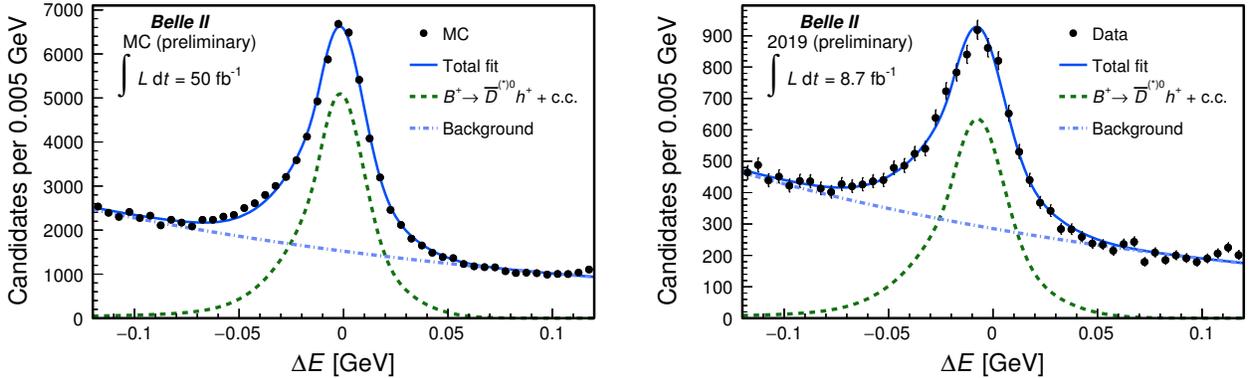


FIG. 3. Fit projection of the maximum likelihood fit to the unbinned distribution of ΔE for $B^+ \rightarrow \bar{D}^{(*)0} h^+$ candidates reconstructed in (left) simulation and (right) data, restricted to $M_{bc} > 5.27 \text{ GeV}/c^2$. The global peak shift and width scaling factor are determined by the fit.

TABLE III. Summary of yields and yields per integrated luminosity obtained from the fit to MC simulation and data. The uncertainties are only statistical.

$B^0 \rightarrow D^{(*)-} h^+$	Yield		Yield/fb ⁻¹	
	MC	Data	MC	Data
Signal	24246 ± 251	4080 ± 114	485 ± 5	469 ± 13
Background	43321 ± 287	7742 ± 129	866 ± 6	890 ± 15

$B^+ \rightarrow \bar{D}^{(*)0} h^+$	MC	Data	MC	Data
	Signal	39706 ± 280	5506 ± 148	794 ± 6
Background	77280 ± 340	14553 ± 176	1546 ± 7	1673 ± 20

6. DATA/MC COMPARISON FOR SIGNAL AND BACKGROUND

We check the data/MC agreement of the flavor tagger output in the fit range by performing an *sPlot* [13] analysis using ΔE as control variable. We determine *sPlot* weights using the fit model developed in Sec. 5. We weight the data with the *sPlot* weights to obtain the individual signal and background distributions in data and compare them with MC simulation. We normalize the simulated samples by scaling the total number of events to those observed in data.

Figures 4 and 5 show the signal and background $q \cdot r$ distributions provided by the category-based flavor tagger for neutral and charged $B \rightarrow D^{(*)} h^+$ candidates. We use the subindex FBDT to label the dilution provided by the flavor tagging algorithm. We compare the signal data distribution with the distribution of correctly associated MC events, and the background data distributions with the distribution of side-band MC events ($M_{bc} < 5.27 \text{ GeV}/c^2$ and same fit range $|\Delta E| < 0.12 \text{ GeV}$). We compare also the signal distributions in data and simulation for the individual tagging categories (Figures 6–8). In general, the results show a good consistency between data and simulation.

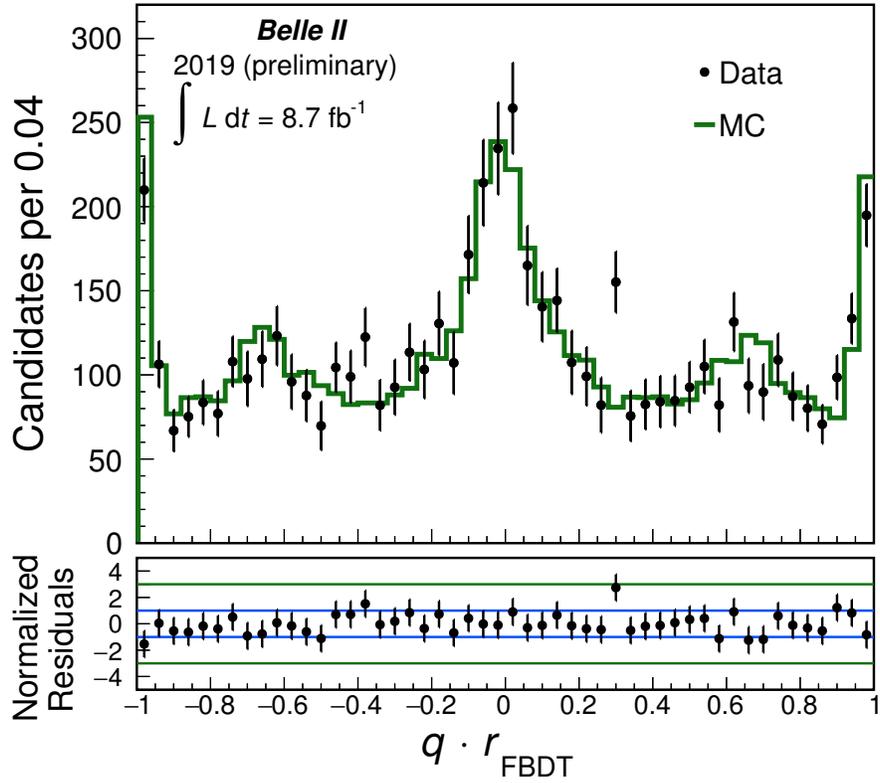
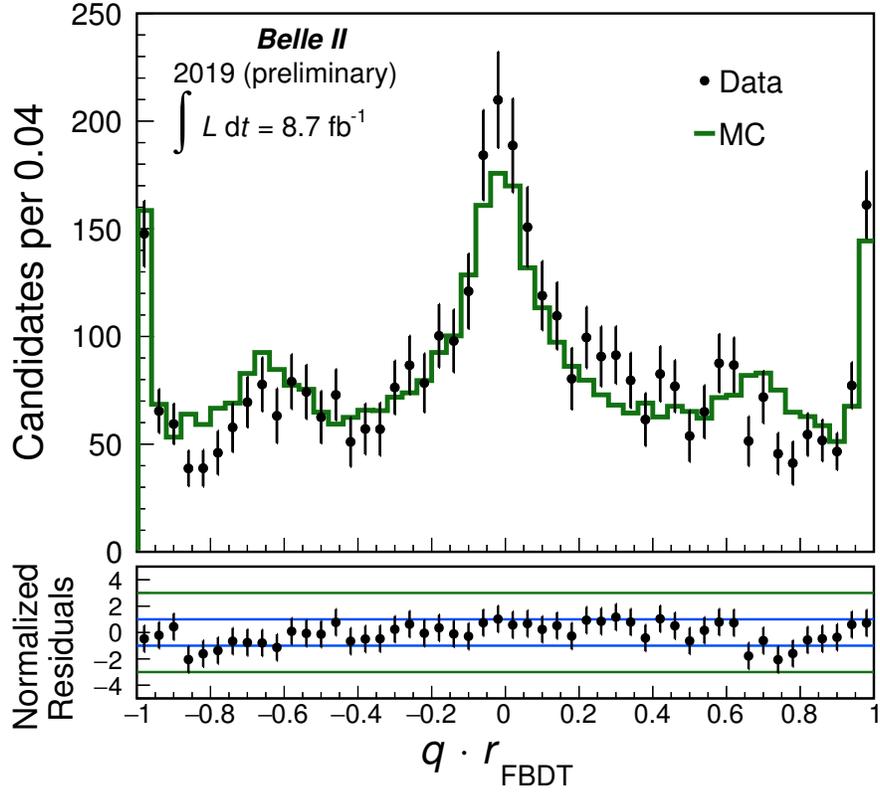


FIG. 4. Normalized $q \cdot r_{\text{FBDT}}$ distributions in data and MC simulation for (top) neutral and (bottom) charged $B \rightarrow D^{(*)}h^+$ candidates. The contribution from the signal component in data is compared with correctly associated signal MC events.

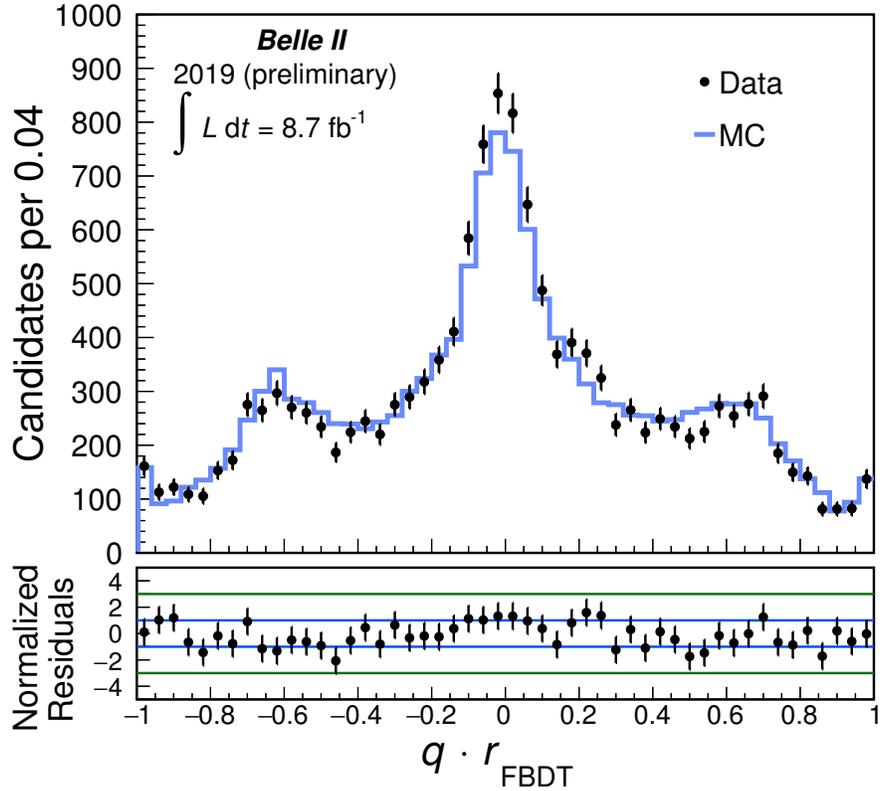
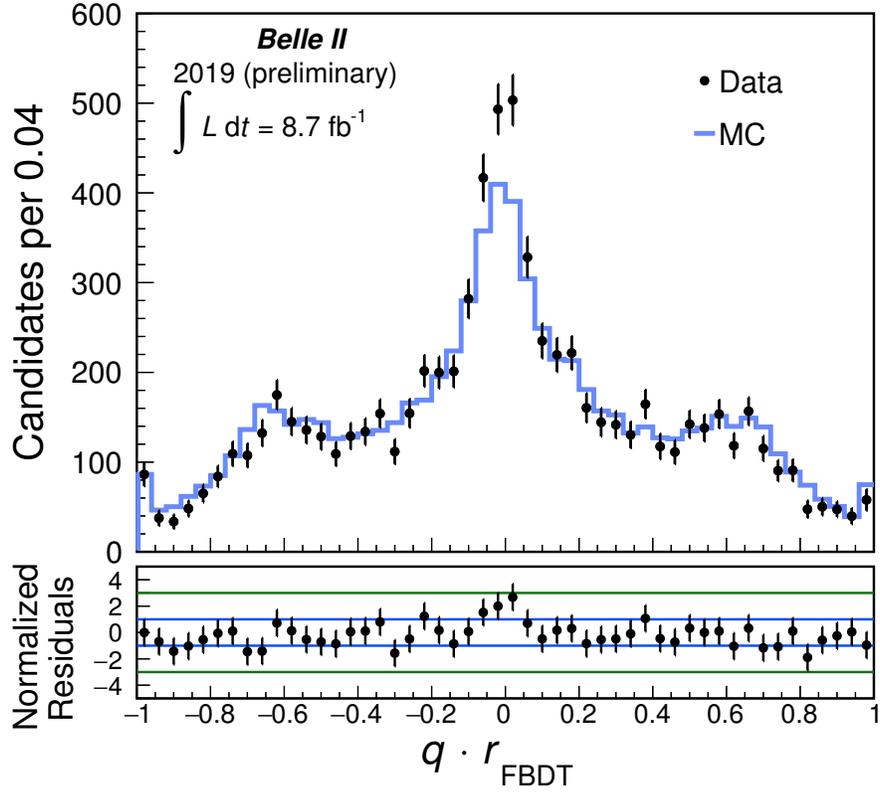


FIG. 5. Normalized $q \cdot r_{\text{FBDT}}$ distributions in data and MC simulation for (top) neutral and (bottom) charged $B \rightarrow D^{(*)}h^+$ candidates. The contribution from the background component in data is compared with simulated events in the side band.

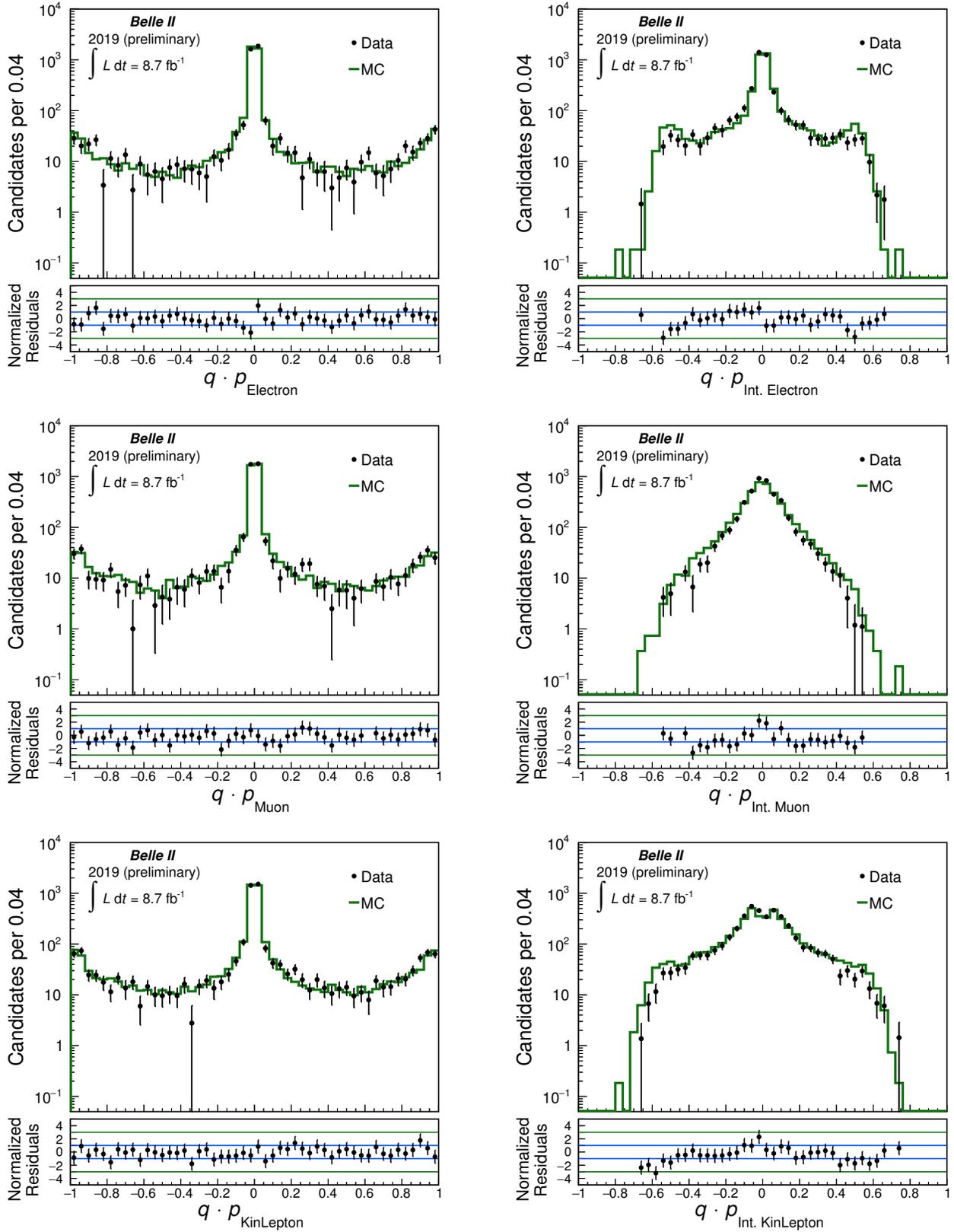


FIG. 6. Normalized output distributions of individual tagging categories in data and MC simulation for $B^0 \rightarrow D^{(*)-} h^+$ candidates. The contribution from the signal component in data is compared with correctly associated signal MC events (1/3).

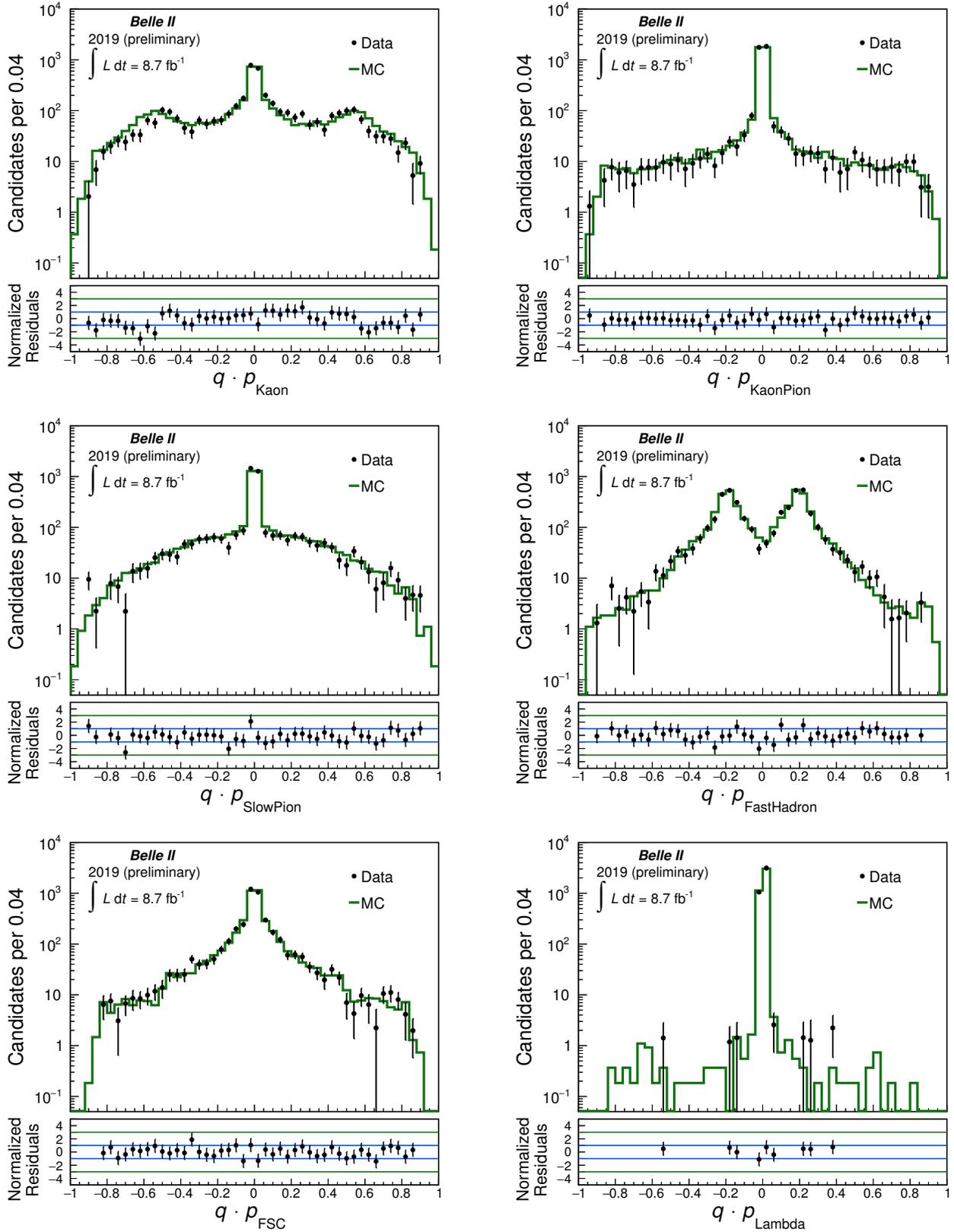


FIG. 7. Normalized output distributions of individual tagging categories in data and MC simulation for $B^0 \rightarrow D^{(*)-}h^+$ candidates. The contribution from the signal component in data is compared with correctly associated signal MC events (2/3).

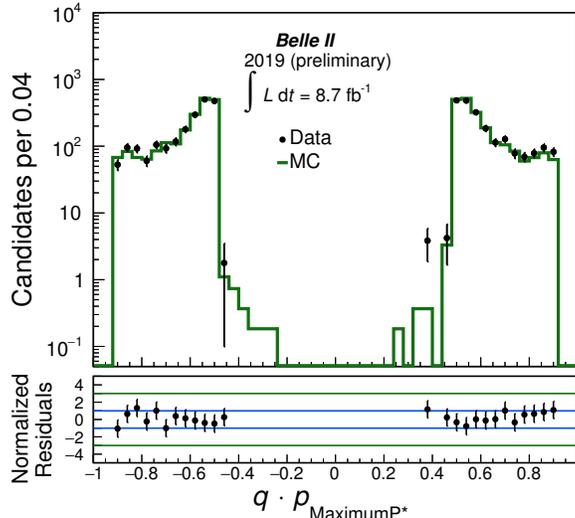


FIG. 8. Normalized output distributions of individual tagging categories in data and MC simulation for $B^0 \rightarrow D^{(*)-} h^+$ candidates. The contribution from the signal component in data is compared with correctly associated signal MC events (3/3).

7. RESULTS

Table IV presents the results for the partial tagging efficiencies and the wrong-tag fractions obtained from the maximum-likelihood fit (Sec. 5) to data. To evaluate the tagging performance, we calculate the total effective efficiency as

$$\varepsilon_{\text{eff}} = \sum_i \varepsilon_{\text{eff},i} = \sum_i \varepsilon_i \cdot (1 - 2w_i)^2,$$

where $\varepsilon_{\text{eff},i}$ is the partial effective efficiency in the i -th r bin. The effective tagging efficiency is a measure for the effective reduction of events due to the flavor dilution r . In CP violation analyses, the statistical uncertainty of measured CP asymmetries is approximately proportional to $1/\sqrt{N_{\text{eff}}} = 1/\sqrt{N \cdot \varepsilon_{\text{eff}}}$, where N_{eff} is the number of effectively tagged events. Thus, one would obtain the same statistical precision for N_{eff} perfectly tagged events or for N events tagged with an effective efficiency ε_{eff} .

We consider systematic uncertainties associated with the model description, the ΔE fit range, the flavor mixing of the background, the fit bias, and the bias introduced by peaking backgrounds.

Model description: we perform pseudo-experiments using an alternative model with a different ΔE parametrization. We perform fits to pseudo-data samples bootstrapped (sampled with replacement) from the generic MC simulation. We fit using the alternative and using the default model and calculate for each fit parameter x_i the difference δx_i between the results obtained with the alternative model and the results obtained with the default model. We obtain the mean difference $\delta \hat{x}_i$ by fitting a Gaussian function to the distribution of δx_i and take the full mean $\delta \hat{x}_i$ as systematic uncertainty.

ΔE Fit range: Figures 2 and 3 show that near the upper limit of the ΔE fit range there is an increase of the background that is slightly above the total fit model. We take into account possible systematic uncertainties due to this slight mismodeling near the upper limit by performing a fit in a reduced range $-0.12 < \Delta E < 0.10$ GeV. We take as systematic

TABLE IV. Results of the maximum-likelihood fit to data: partial tagging efficiencies, wrong-tag fractions, partial effective efficiencies and total effective efficiency for neutral and charged B candidates. The results are given with statistical and systematic uncertainties in percent.

$B^0 \rightarrow D^{(*)-} h^+$			
r - Interval	ε_i	w_i	$\varepsilon_{\text{eff},i}$
0.000 – 0.100	$20.3 \pm 1.8 \pm 0.3$	$47.4 \pm 4.1 \pm 0.9$	$0.1 \pm 0.2 \pm 0.1$
0.100 – 0.250	$17.4 \pm 0.9 \pm 0.1$	$42.8 \pm 4.4 \pm 0.5$	$0.4 \pm 0.4 \pm 0.1$
0.250 – 0.500	$21.2 \pm 0.9 \pm 0.4$	$26.9 \pm 3.7 \pm 0.1$	$4.5 \pm 1.5 \pm 0.1$
0.500 – 0.625	$11.1 \pm 0.7 \pm 0.2$	$16.7 \pm 5.0 \pm 2.4$	$4.9 \pm 1.5 \pm 0.7$
0.625 – 0.750	$9.6 \pm 0.7 \pm 0.5$	$9.2 \pm 5.1 \pm 4.0$	$6.4 \pm 1.7 \pm 1.3$
0.750 – 0.875	$7.0 \pm 0.6 \pm 0.2$	$12.0 \pm 5.6 \pm 0.8$	$4.1 \pm 1.2 \pm 0.2$
0.875 – 1.000	$13.4 \pm 0.7 \pm 0.3$	$0.0 \pm 3.3 \pm 0.1$	$13.4 \pm 1.9 \pm 0.3$
Total	$\varepsilon_{\text{eff}} = \sum_i \varepsilon_i \cdot (1 - 2w_i)^2 = 33.8 \pm 3.6 \pm 1.6$		

$B^+ \rightarrow \bar{D}^{(*)0} h^+$			
r - Interval	ε_i	w_i	$\varepsilon_{\text{eff},i}$
0.000 – 0.100	$17.7 \pm 1.7 \pm 0.4$	$46.5 \pm 2.7 \pm 0.4$	$0.1 \pm 0.1 \pm 0.1$
0.100 – 0.250	$16.0 \pm 0.8 \pm 0.2$	$41.6 \pm 2.7 \pm 1.6$	$0.5 \pm 0.3 \pm 0.2$
0.250 – 0.500	$21.3 \pm 0.9 \pm 0.1$	$29.6 \pm 2.1 \pm 0.9$	$3.6 \pm 0.8 \pm 0.3$
0.500 – 0.625	$10.8 \pm 0.7 \pm 0.2$	$13.5 \pm 2.6 \pm 0.8$	$5.8 \pm 0.9 \pm 0.3$
0.625 – 0.750	$10.6 \pm 0.7 \pm 0.5$	$11.0 \pm 2.3 \pm 0.7$	$6.5 \pm 0.9 \pm 0.4$
0.750 – 0.875	$9.1 \pm 0.6 \pm 0.1$	$5.6 \pm 1.8 \pm 0.2$	$7.2 \pm 0.7 \pm 0.1$
0.875 – 1.000	$14.5 \pm 0.6 \pm 0.4$	$2.8 \pm 0.8 \pm 0.3$	$12.9 \pm 0.7 \pm 0.4$
Total	$\varepsilon_{\text{eff}} = \sum_i \varepsilon_i \cdot (1 - 2w_i)^2 = 36.6 \pm 1.8 \pm 0.7$		

uncertainty for each fit parameter x_i the difference between the results obtained in the reduced range and the results obtained in the default ΔE range.

Background mixing: our fit takes into account the uncertainty on the world average for the signal χ_d in the Gaussian constraint. However, we assume that there is no mixing in the background ($\chi_d^{\text{bkg}} = 0$). Since the background includes $B^0 \bar{B}^0$ events, we study the effect of flavor mixing in the background by varying the value of the background χ_d^{bkg} by a small amount $\pm \delta \chi_d^{\text{bkg}}$, corresponding to the statistical uncertainty when we leave χ_d^{bkg} free to float. We then take for each fit parameter x_i half the difference between the results for $\chi_d^{\text{bkg}} + \delta \chi_d^{\text{bkg}}$ and for $\chi_d^{\text{bkg}} - \delta \chi_d^{\text{bkg}}$ as systematic uncertainty.

Fit bias: for each fit parameter x_i , we determine the fit bias using the residuals from bootstrapped pseudo-experiments. The residuals are the differences between the fit results for the individual pseudo-data samples and the fit results for the parent MC sample. We take the full bias as systematic uncertainty.

Peaking background bias: we consider the bias caused by the peaking background, which is not included in the fit model, by calculating the difference between the results of

the fit to the full MC sample and the true values determined using MC information. We take the full difference as systematic uncertainty.

We find the systematic uncertainty associated with the peaking background bias to be the dominant one around 40% of the statistical uncertainty, followed by the model description around 6% and the fit bias around 3%. The systematic uncertainties due to the fit range and due to the background mixing are around or below 1% of the statistical uncertainty and therefore negligible. In future calibrations using larger data samples, we will consider the peaking background in the fit model and thus we expect the associated systematic uncertainty to decrease. With larger samples, we also expect to improve the fit model description of the data and thus to reduce the uncertainty due to the model description.

8. LINEARITY CHECK

By definition, the dilution factor r is equal to $1 - 2w$. We probe if the dilution r provided by the flavor tagger corresponds to the actual definition by performing a linearity check. Figure 9 shows the linearity check for simulation and data. For simulation, we determine the true wrong-tag fraction w_{MC} by comparing the MC truth with the flavor tagger output, and calculate the true dilution $r_{\text{MC}} = 1 - 2w_{\text{MC}}$. The mean dilution $\langle r_{\text{FBDT}} \rangle$ is simply the mean of $|q \cdot r_{\text{FBDT}}|$ for correctly associated MC events in each r bin. For data, we obtain the mean $\langle r_{\text{FBDT}} \rangle = \langle |q \cdot r_{\text{FBDT}}| \rangle$ values from the signal $q \cdot r_{\text{FBDT}}$ distribution provided by the *sPlot* analysis in Sec. 6. The dilution $r = 1 - 2 \cdot w$ in data is obtained from the fit results for w . The linearity verifies the equivalence in average between the dilution provided by the flavor tagger and the measured one within the uncertainties. For charged B candidates, we observe a slightly non-linear behaviour which we attribute to the fact that the flavor tagger is optimized and trained only for neutral B mesons. However, we observe a good agreement between data and simulation for both neutral and charged B candidates.

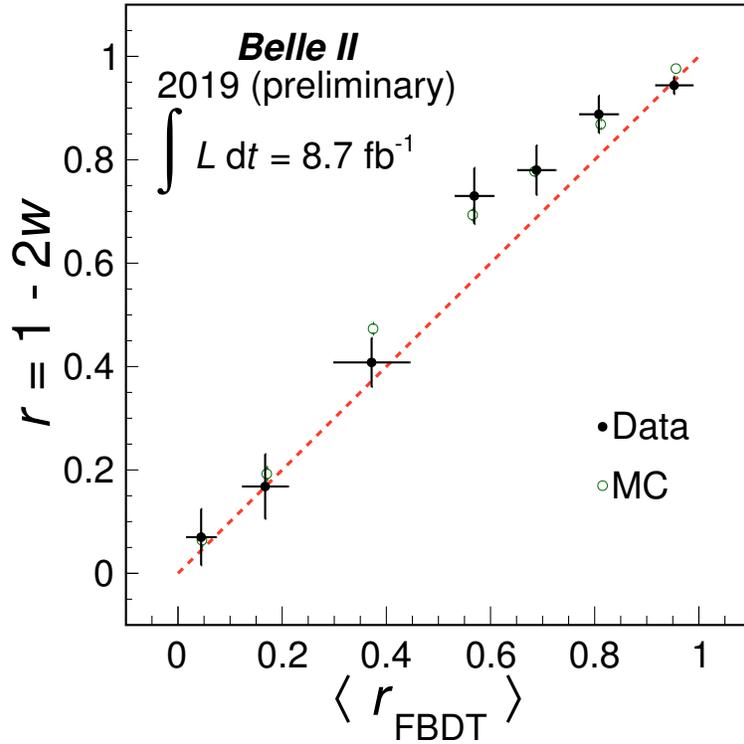
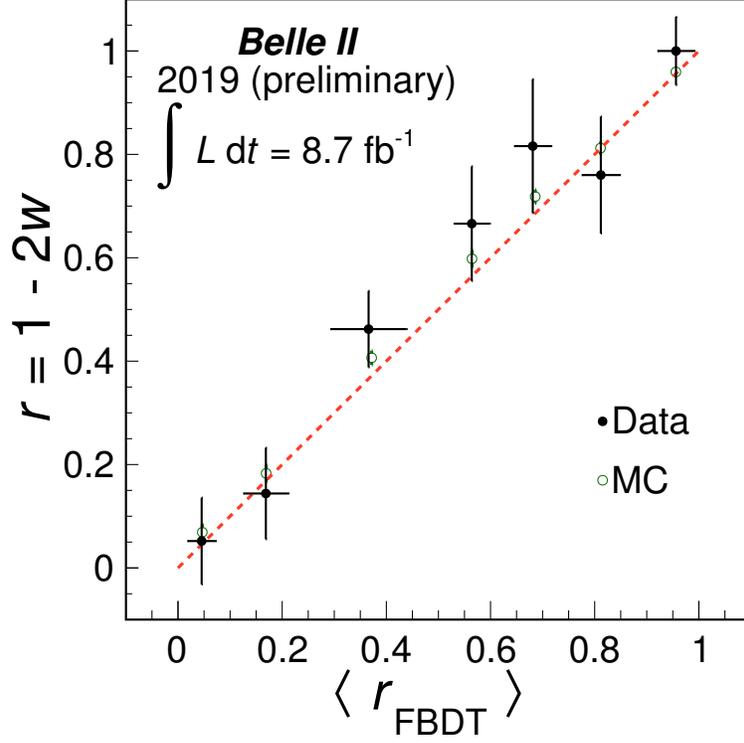


FIG. 9. Dilution factor $r = 1 - 2w$ as a function of the mean dilution $\langle |q \cdot r_{\text{FBDT}}| \rangle$ provided by the flavor tagger in data and MC simulation for (top) neutral and (bottom) charged $B \rightarrow D^{(*)}h^+$ candidates. The red guidelines correspond to a linear function with an intercept at 0 and a slope of 1, i.e. to a perfect agreement between predicted and measured dilution.

9. COMPARISON WITH BELLE

Comparison of the current results with Belle’s latest results [10] on flavor tagging provides interesting insight to assess Belle II’s current and projected performance. We compare the wrong-tag fractions and the efficiencies in each r -bin, and the total effective efficiencies, which are shown in Table V and Fig. 10. The Belle flavor tagger reached an effective efficiency of $(30.1 \pm 0.4)\%$ on Belle data [10]. In comparison with the previous Belle algorithm, the new Belle II category-based flavor tagger considers more flavor signatures and more input variables, and is based on multivariate methods avoiding cut-based approaches.

TABLE V. Partial efficiencies ε_i and wrong-tag fractions w_i obtained with the Belle II flavor tagger in 2019 Belle II data and with the Belle flavor tagger in Belle data [10] taken with the second silicon-vertex detector configuration (SVD2). Statistical and systematical uncertainties are added in quadrature. All values are given in percent.

$B^0 \rightarrow D^{(*)-} h^+$	$\varepsilon_i \pm \delta\varepsilon_i$		$w_i \pm \delta w_i$		$\varepsilon_{\text{eff},i} \pm \delta\varepsilon_{\text{eff},i}$	
	Belle II	Belle	Belle II	Belle	Belle II	Belle
0.000 – 0.100	20.3 ± 1.8	22.2 ± 0.4	47.4 ± 4.2	50.0	0.1 ± 0.2	0.0
0.100 – 0.250	17.4 ± 0.9	14.5 ± 0.3	42.8 ± 4.4	41.9 ± 0.4	0.4 ± 0.4	0.4 ± 0.1
0.250 – 0.500	21.2 ± 1.0	17.7 ± 0.4	26.9 ± 3.7	31.9 ± 0.3	4.5 ± 1.5	2.3 ± 0.1
0.500 – 0.625	11.1 ± 0.7	11.5 ± 0.3	16.7 ± 5.5	22.3 ± 0.4	4.9 ± 1.7	3.5 ± 0.1
0.625 – 0.750	9.6 ± 0.9	10.2 ± 0.3	9.2 ± 6.5	16.3 ± 0.4	6.4 ± 2.1	4.6 ± 0.2
0.750 – 0.875	7.0 ± 0.6	8.7 ± 0.3	1.2 ± 5.7	10.4 ± 0.4	4.0 ± 1.2	5.5 ± 0.1
0.875 – 1.000	13.4 ± 0.8	15.3 ± 0.3	0.0 ± 3.3	2.5 ± 0.3	13.4 ± 1.9	13.8 ± 0.3
	Total	$\varepsilon_{\text{eff}} = \sum_i \varepsilon_i \cdot (1 - 2w_i)^2 = 33.8 \pm 3.9$			30.1 ± 0.4	

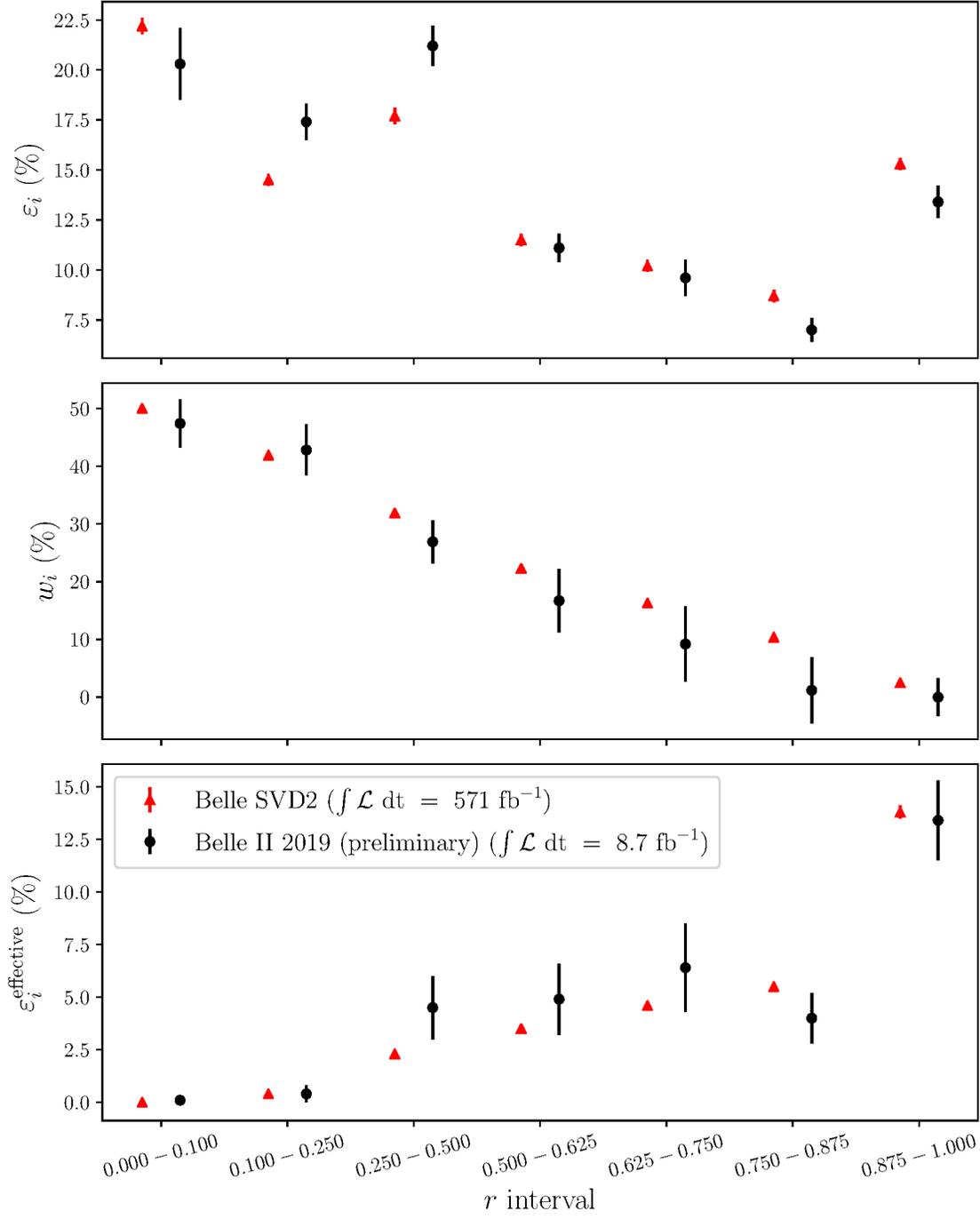


FIG. 10. Performance of the Belle II flavor tagger in 2019 Belle II data and of the Belle flavor tagger in Belle data [10] taken with the second silicon-vertex detector configuration (SVD2).

10. SUMMARY

We report on the first calibration of the standard Belle II B -flavor tagger using 2019 Belle II data. The ΔE distributions of reconstructed charmed B candidates, restricted in M_{bc} , are fit to identify the B signals and measure the tagging efficiencies and the fractions of wrongly tagged events from the flavor evolution of the signal $B\bar{B}$ pairs in a time-integrated way. The total effective efficiency for neutral B candidates is measured to be

$$\varepsilon_{\text{eff}} = (33.8 \pm 3.6(\text{stat}) \pm 1.6(\text{sys}))\%,$$

and for charged B candidates

$$\varepsilon_{\text{eff}} = (36.6 \pm 1.8(\text{stat}) \pm 0.7(\text{sys}))\%.$$

The performance of the flavor tagger is generally compatible with expectations from simulation (Fig. 9), establishing a good understanding of the detector performance. The performance is also comparable with the best one obtained by the Belle experiment within the uncertainties (Fig. 10). This work marks a first milestone for future calibrations which will play an essential role in measurements of CP -asymmetries.

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